

# An introduction to extreme value statistics (part 1/2)

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Dresden  
October 05, 2015

## Part 1 – iid/weakly correlated variables

- ▶ Maxima, exceedances.
- ▶ iid theory, extreme value distributions.
- ▶ Point processes, Poisson approximation.
- ▶ Extensions to weakly correlated stationary sequences.
- ▶ Modeling examples.

## Part 2 – examples from statistical physics

- ▶ Remarks on applications statistical physics.
- ▶ DNA replication times.
- ▶ Order parameter in percolation.
- ▶ Minimum path in tree-like structures.
- ▶ Maxima of some Gaussian processes.
- ▶ Integer partitions and the ideal Bose gas.

# References

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Theory

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Theory

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Lecture notes

# Outline

## EVS and Society

## GEV & GPD

- Block maxima & exceedances

- Limiting distributions

- Domains of attraction

## Point process perspective on extremes

## Weakly correlated stationary sequences

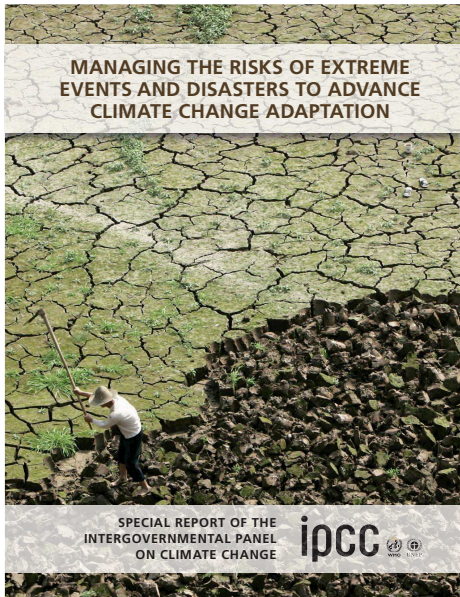
- Asymptotic independence of maxima

- Clustering & extremal index

## Modelling examples

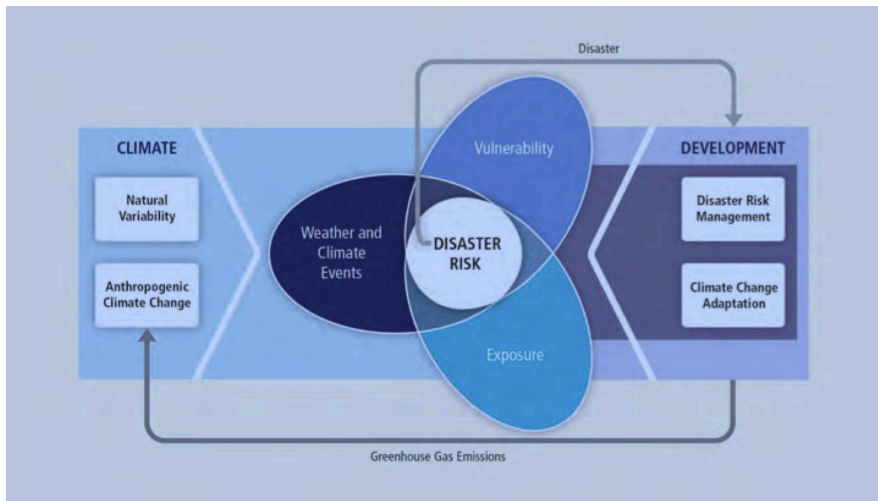
# EVS and Climate Change

IPCC 2012



# EVS and Climate Change

"Managing the Risks of Extreme Events and Disasters to Advance Climate Change Adaptation"



# 15th century IPCC

Behringer 1999; Summers 2003



Figure 3. Witches cause a hailstorm. Illustration from the title page of Ulrich Meitner's *De laniis et phitonicis mulieribus* [Concerning Witches and Sorceresses], Cologne, 1489.

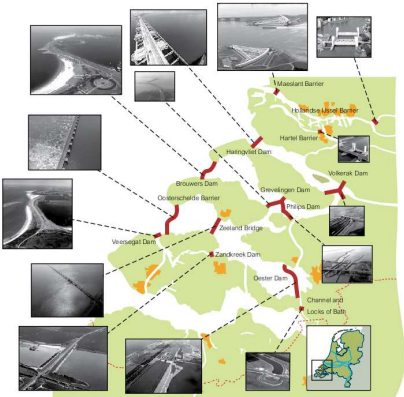


"...in some parts of Northern Germany, as well as in the provinces, townships, territories, districts, and dioceses of Mainz, Cologne, Trèves, Salzburg, and Bremen, many persons... by their incantations, spells, conjurations, and other accursed charms and crafts, enormities and horrid offences, ...have blasted the produce of the earth, the grapes of the vine, the fruits of the trees,... vineyards, orchards, meadows, pasture-land, corn, wheat, and all other cereals..."

(Pope Innocent VIII, 1484)



# Deltawerken



# Outline

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## GEV & GPD

Block maxima & exceedances

Limiting distributions

Domains of attraction

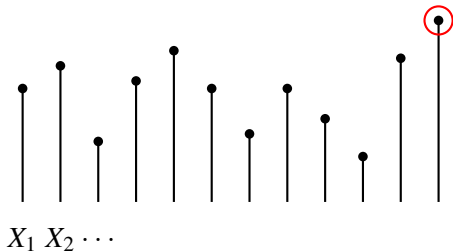
Point process perspective on extremes

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Asymptotic independence of maxima

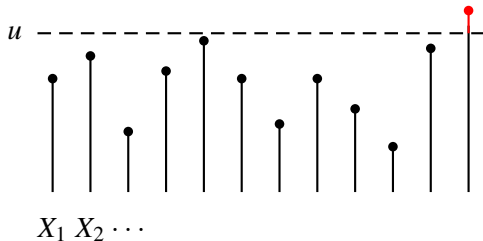
Clustering & extremal index

Modelling examples



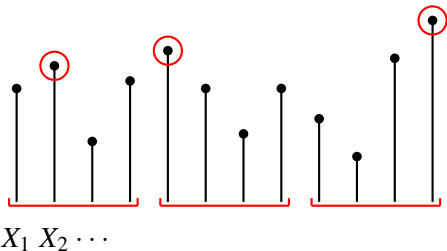
$$\mathbb{P}(\max\{X_1, \dots, X_n\} \leq x)$$

**GEV**  
(Generalised Extreme Value)



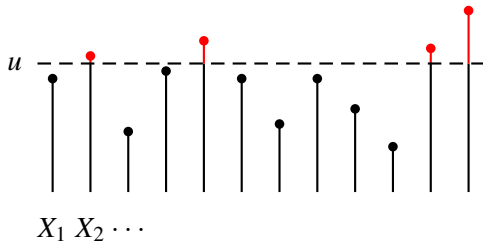
$$\mathbb{P}(X \leq u + y | X > u)$$

**GPD**  
(Generalised Pareto)



$$\mathbb{P}(\max\{X_1, \dots, X_n\} \leq x)$$

**GEV**  
(Generalised Extreme Value)



$$\mathbb{P}(X \leq u + y | X > u)$$

**GPD**  
(Generalised Pareto)

# Statement of the problem

Draw  $n$  **independent** and **identically** distributed random variables  $X_1, \dots, X_n$  from a parent distribution  $F(x)$ .

What is the distribution of the maximum

$$M_n = \max\{X_1, \dots, X_n\} \quad ?$$

For iid random variables

$$\begin{aligned}\mathbb{P}(M_n \leq x) &= \mathbb{P}(X_1 \leq x, \dots, X_n \leq x) \\ &= \mathbb{P}(X_1 \leq x) \cdots \mathbb{P}(X_n \leq x) \\ &= F^n(x).\end{aligned}$$

# Limit distribution for maximum

## Affine rescaling

E.g. for uniform distribution  $X_i \sim U(0, 1)$

$$F^n(x) = x^n \xrightarrow{n \rightarrow \infty} \begin{cases} 1 & x = 1 \\ 0 & x < 1. \end{cases}$$

Can we do better than 0-1 behaviour?

Choose  $a_n$  (scale) and  $b_n$  (location) such that

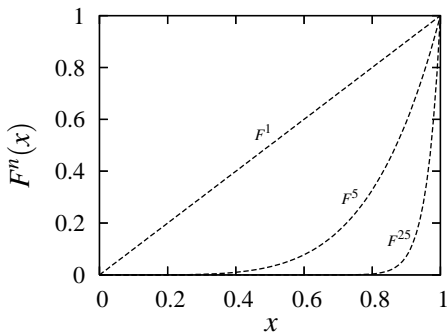
$$\lim_{n \rightarrow \infty} F^n(a_n x + b_n) = G(x)$$

is a non-degenerate limiting distribution.

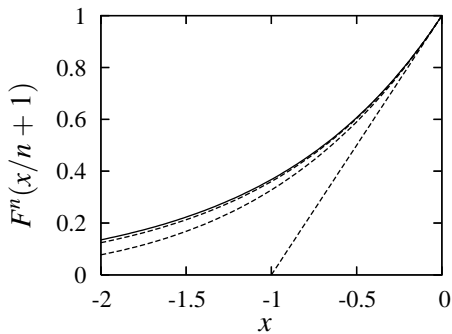
# Example: uniform distribution

$a_n$  (scale),  $b_n$  (location)

$$F(x) = x, \quad x \in [0, 1]$$



$$a_n = 1/n, b_n = 1$$



## Example: uniform distribution

$$F(x) = x$$

With  $a_n = 1/n, b_n = 1$ :

$$\lim_{n \rightarrow \infty} F^n(a_n x + b_n) = \lim_{n \rightarrow \infty} \left( \frac{x}{n} + 1 \right)^n = \exp(x),$$

where

$$G(x) = \exp(x), \quad x \in (-\infty, 0]$$

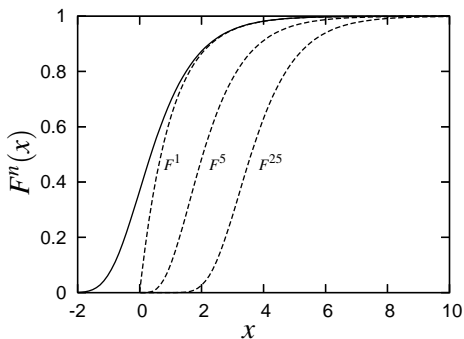
is a **Weibull** distribution.



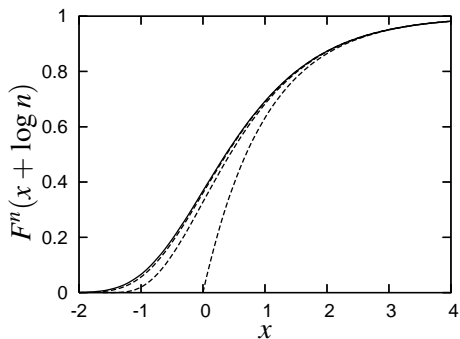
# Example: exponential distribution

$a_n$  (scale),  $b_n$  (location)

$$F(x) = 1 - \exp(-x), \quad x \in [0, \infty)$$



$$a_n = 1, b_n = \log n$$



## Example: exponential distribution

$$F(x) = 1 - \exp(-x)$$

With  $a_n = 1$ ,  $b_n = \log n$ :

$$\begin{aligned} \lim_{n \rightarrow \infty} \{1 - \exp[-(x + \log n)]\}^n &= \lim_{n \rightarrow \infty} \left[1 - \frac{\exp(-x)}{n}\right]^n \\ &= \exp[-\exp(-x)], \end{aligned}$$

where

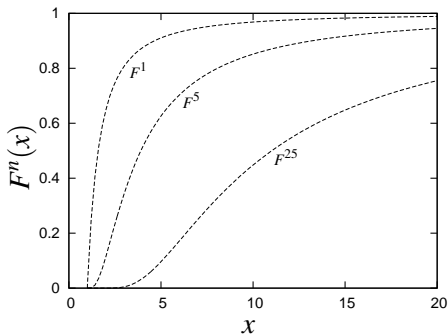
$$G(x) = \exp[-\exp(-x)], \quad x \in (-\infty, \infty)$$

is the **Gumbel** distribution.

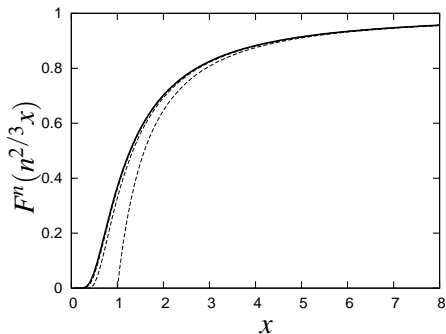
# Example: a Pareto distribution

$a_n$  (scale),  $b_n$  (location)

$$F(x) = 1 - x^{-3/2}, \quad x \in [1, \infty)$$



$$a_n = n^{2/3}, b_n = 0$$



## Example: a Pareto distribution

$$F(x) = 1 - x^{-3/2}$$

With  $a_n = n^{2/3}$ ,  $b_n = 0$ :

$$\begin{aligned}\lim_{n \rightarrow \infty} \{1 - (n^{2/3}x)^{-3/2}\}^n &= \lim_{n \rightarrow \infty} \left[1 - \frac{x^{-3/2}}{n}\right]^n \\ &= \exp[-x^{-3/2}],\end{aligned}$$

where

$$G(x) = \exp[-x^{-3/2}], \quad x \in [0, \infty)$$

is a **Fréchet** distribution.

# Generalised extreme value distribution

## Jenkinson-von Mises representation

Overall maximum = maximum of block maxima:

$$\max \left\{ \underbrace{\{X_{i_1}, \dots, X_{i_n}\}}_{F^{mn}(x)}, \dots, \underbrace{\{X_{j_1}, \dots, X_{j_n}\}}_{G^m(a_n^{-1}(x-b_n))} \right\} = \max \left\{ \underbrace{M_{n;1}, \dots, M_{n;m}}_{G^m(a_n^{-1}(x-b_n))} \right\}$$

$$G^m(a_n^{-1}a_{mn}x + a_n^{-1}(b_{mn} - b_n)) \rightarrow G^m(a_mx + b_m) = G(x)$$

GEV distributions are **max-stable**:

$$G^m(a_mx + b_m) = G(x).$$

Solve functional equation:

$$G_\xi(x) = \exp[-(1 + \xi x)^{-1/\xi}], \quad 1 + \xi x > 0.$$

# Generalised extreme value distribution

Fisher, Tippett 1928

Weibull,  $\xi < 0$

$$G_{\xi < 0}(x) = \begin{cases} \exp(-(-x)^{-1/\xi}), & x \leq 0 \\ 1, & x > 0 \end{cases}$$

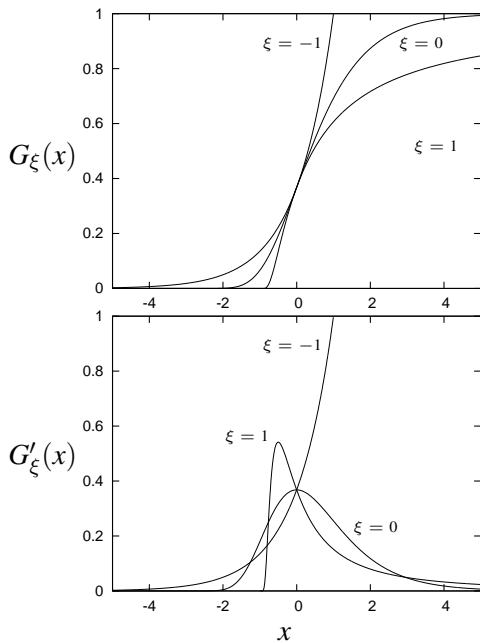
Gumbel,  $\xi = 0$

$$G_{\xi = 0}(x) = \exp(-\exp(-x)), \quad -\infty < x < \infty$$

Fréchet,  $\xi > 0$

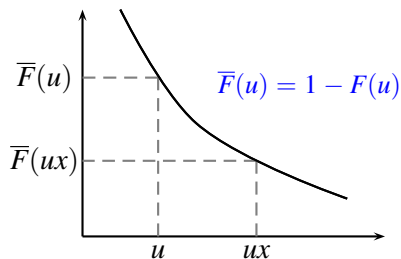
$$G_{\xi > 0}(x) = \begin{cases} 0, & x \leq 0 \\ \exp(-x^{-1/\xi}), & x > 0 \end{cases}$$

# Generalised extreme value distribution



# Characterising tail behaviour

## Regular variation



$\bar{F} \in \mathcal{R}_{-\alpha}$  (regularly varying with exponent  $-\alpha$ ) if

$$\lim_{u \rightarrow \infty} \frac{\bar{F}(ux)}{\bar{F}(u)} = x^{-\alpha}, \quad x > 0.$$



## Domains of attraction: Fréchet

E.g. for power law  $F(u) = 1 - u^{-1/\xi}$ ,  $u \in [1, \infty)$ ,

$$\begin{aligned}\lim_{u \rightarrow \infty} \frac{\overline{F}(ux)}{\overline{F}(u)} &= \lim_{u \rightarrow \infty} \frac{(ux)^{-1/\xi}}{u^{-1/\xi}} \\ &= x^{-1/\xi}.\end{aligned}$$

## Domains of attraction: Weibull

If  $x_F < \infty$ , check behaviour of  $\bar{F}(x_F - u)$  as  $u \downarrow 0$ .

E.g. for uniform  $F(u) = u, u \in [0, 1]$ ,

$$\begin{aligned}\lim_{u \downarrow 0} \frac{\bar{F}(x_F - ux)}{\bar{F}(x_F - u)} &= \lim_{u \downarrow 0} \frac{ux}{u} \\ &= x.\end{aligned}$$

## Domains of attraction: Weibull

E.g. for truncated exponential

$$F(u) = \frac{1 - \exp(-u)}{1 - \exp(-x_F)}, \quad u \in [0, x_F],$$

$$\begin{aligned} \lim_{u \downarrow 0} \frac{\bar{F}(x_F - ux)}{\bar{F}(x_F - u)} &= \lim_{u \downarrow 0} \frac{\exp(ux) - 1}{\exp(u) - 1} \\ &= x. \end{aligned}$$

## Domains of attraction: Gumbel

E.g. for exponential,  $F(u) = 1 - \exp(-u)$ ,  $u \in [0, \infty)$ ,

$$\begin{aligned}\lim_{u \rightarrow \infty} \frac{\bar{F}(ux)}{\bar{F}(u)} &= \lim_{u \rightarrow \infty} \exp(-u(x-1)) \\ &= \begin{cases} 0, & x > 1 \\ \infty, & x < 1. \end{cases}\end{aligned}$$

$\exp(-u) \in \mathcal{R}_{-\infty}$  (rapid variation).

## Domains of attraction $\mathcal{D}(\mathcal{G}_\xi)$

$\bar{F} \in \mathcal{D}(\mathcal{G}_\xi)$  iff for some  $g > 0$

$$\lim_{u \uparrow x_F} \frac{\bar{F}(u + xg(u))}{\bar{F}(u)} = (1 + \xi x)^{-1/\xi}, \quad 1 + \xi x > 0,$$

where  $g$  (non-unique) can be taken to be

$$g(u) = \begin{cases} \xi u, & \xi > 0 \text{ (Fréchet)}, \\ -\xi(x_F - u), & \xi < 0 \text{ (Weibull)}, \\ \int_u^{x_F} dt \bar{F}(t)/\bar{F}(u), & \xi = 0 \text{ (Gumbel)}. \end{cases}$$

Norming constants (non-unique) can be taken to be

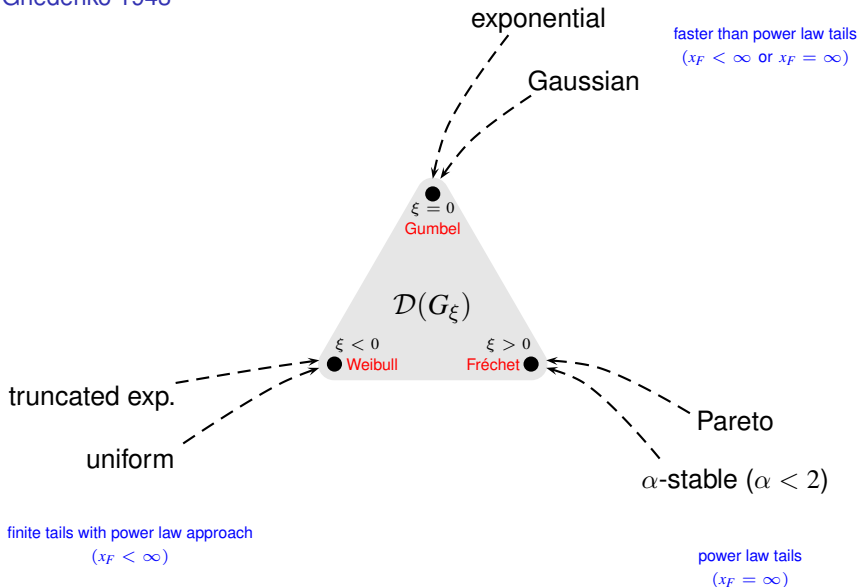
$$a_n = F^{-1}(1 - n^{-1}) \quad b_n = 0 \quad \xi > 0 \text{ (Fréchet)},$$

$$a_n = x_F - F^{-1}(1 - n^{-1}) \quad b_n = x_F \quad \xi < 0 \text{ (Weibull)},$$

$$a_n = g(b_n) \quad b_n = F^{-1}(1 - n^{-1}) \quad \xi = 0 \text{ (Gumbel)}.$$

# Domains of attraction $\mathcal{D}(\mathcal{G}_\xi)$

Gnedenko 1943



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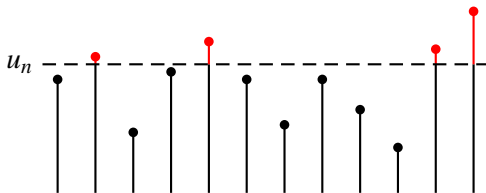
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# Poisson approximation

Number of exceedances  $n_u$  is a Poisson random variable



## Poisson process

$$\begin{aligned}n_u &\stackrel{d}{\sim} \text{Bin}(n, \bar{F}(u_n)) \\ \mathbb{E}[n_u] &= n\bar{F}(u_n) \rightarrow \tau \\ n_u &\rightarrow \text{Poi}(\tau)\end{aligned}$$

$$\text{Poisson approx. to extremes} \begin{cases} n\bar{F}(u_n) \rightarrow \tau \\ \mathbb{P}(M_n \leq u_n) = \mathbb{P}(n_u = 0) \rightarrow \exp(-\tau) \end{cases}$$

Generalises for  $k$ th maximum:

$$\mathbb{P}(M_n^{(k)} \leq u_n) = \mathbb{P}(n_u < k) \rightarrow \exp(-\tau) \sum_{m=1}^{k-1} \frac{\tau^m}{m!}.$$



# Consistency with GEV

## Refining the Poisson approximation

Since

$$\lim_{n \rightarrow \infty} F^n(a_n x + b_n) = G_\xi(x),$$

$$n \log F(a_n x + b_n) \rightarrow \log G_\xi(x)$$

$$n \log(1 - \bar{F}(a_n x + b_n)) \rightarrow \log G_\xi(x)$$

$$n\bar{F}(a_n x + b_n) \approx -\log G_\xi(x).$$

With the choice  $u_n = a_n x + b_n$ ,

$$\begin{aligned}\tau &= -\log G_\xi(x) \\ &= (1 + \xi x)^{-1/\xi},\end{aligned}$$

so that

$$\mathbb{P}(M_n \leq a_n x + b_n) = \exp(-(1 + \xi x)^{-1/\xi}).$$

## Connection between GEV and GPD

From  $n\bar{F}(a_n x + b_n) \approx -\log G_\xi(x)$ ,

$$\bar{F}(x) \approx \frac{1}{n} \left[ 1 + \xi \frac{(x - b_n)}{a_n} \right]^{-1/\xi},$$

so excess  $y$  over threshold  $u$  is distributed as

$$\begin{aligned} \mathbb{P}(X > u + y | X > u) &= \frac{\bar{F}(u + y)}{\bar{F}(u)} = \frac{\left[ 1 + \xi \frac{(u + y - b_n)}{a_n} \right]^{-1/\xi}}{\left[ 1 + \xi \frac{(u - b_n)}{a_n} \right]^{-1/\xi}} \\ &= \left( 1 + \frac{\xi}{\beta} y \right)^{-1/\xi} \quad (\text{GPD}), \end{aligned}$$

with

$$\beta = a_n + \xi(u - b_n).$$

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**Weakly correlated stationary sequences**

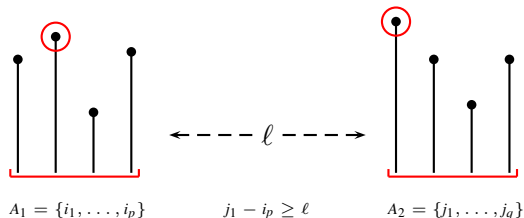
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# Asymptotic independence of maxima

Loyes 1965; Leadbetter, Lindgren, Rootzén 1983



Stationary sequence  $X_n$  satisfies  $D(u_n)$  (mixing) condition if

$$\left| \mathbb{P} \left( \max_{i \in A_1 \cup A_2} X_i \leq u_n \right) - \mathbb{P} \left( \max_{i \in A_1} X_i \leq u_n \right) \mathbb{P} \left( \max_{i \in A_2} X_i \leq u_n \right) \right| \leq \alpha_{n,\ell},$$

where  $\alpha_{n,\ell} \rightarrow 0$  as  $n \rightarrow \infty$  for some  $\ell = \ell_n = o(n)$ .

# Berman's condition

Berman 1964

Stationary Gaussian sequence satisfies  $D(u_n)$  if its autocorrelation function

$$\lim_{n \rightarrow \infty} r_n \log n = 0.$$

- ▶ AR(1) process:  $r_n \sim \exp(-n)$ .
- ▶ Long-range dependence:  $r_n \sim n^{-\gamma}, \gamma \in (0, 1)$ .

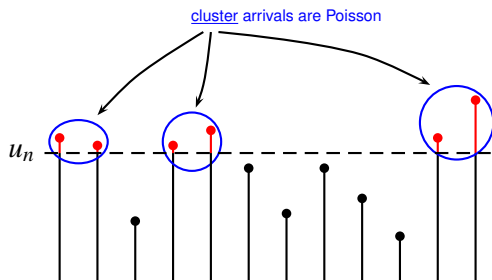
**GEV/GPD limit distributions still apply!**

Counterexample:

- ▶ Random walk:  $r_n \sim n$  (Berman's condition fails).

# Clustering & extremal index

Compound Poisson point process



$$n\bar{F}(u_n) \rightarrow \tau$$

$$\mathbb{P}(M_n \leq u_n) \rightarrow \exp(-\theta\tau)$$

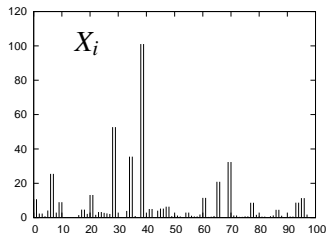
$$\text{cluster size} \approx 1/\theta$$

Heuristic interpretation of  $\theta \in [0, 1]$ :

- ▶ Inverse mean cluster size.
- ▶ Measure of loss of iid degrees of freedom.
- ▶ Multiplicity of associated compound Poisson process.

Can check for  $\theta < 1$  via an **anticlustering condition**  $D'(u_n)$ .

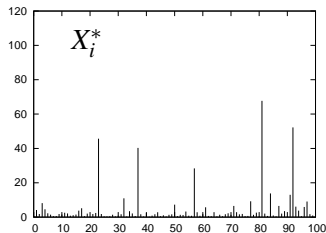
## Example: $X_i = \max\{Y_i, Y_{i+1}\}$



E.g.  $F_Y(y) = \exp(-1/2y), y \in [0, \infty)$

$$\begin{aligned}\mathbb{P}(X_i \leq x) &= \mathbb{P}(Y_i \leq x, Y_{i+1} \leq x) \\ &= \exp(-1/x)\end{aligned}$$

$$F_X(x) = \exp(-1/x) \quad \text{(dependent)}$$



$$F_{X^*}(x) = \exp(-1/x) \quad \text{(iid)}$$

Example:  $X_i = \max\{Y_i, Y_{i+1}\}$

For iid sequence:

$$\begin{aligned}\mathbb{P}(M_n^* \leq nx) &= F_{X^*}^n(nx) \\ &= \exp(-1/x).\end{aligned}$$

For dependent sequence:

$$\begin{aligned}\mathbb{P}(M_n \leq nx) &= \mathbb{P}(X_1 \leq nx, \dots, X_n \leq nx) \\ &= \mathbb{P}(Y_1 \leq nx, Y_2 \leq nx, \dots, Y_n \leq nx, Y_{n+1} \leq nx) \\ &= F_Y^{n+1}(nx) \\ &\rightarrow \exp(-1/x)^{1/2}, \quad n \rightarrow \infty.\end{aligned}$$

So

$$G_X(x) = G_{X^*}^\theta(x),$$

with extremal index  $\theta = 1/2$ .



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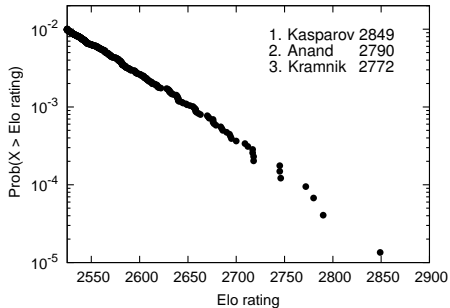
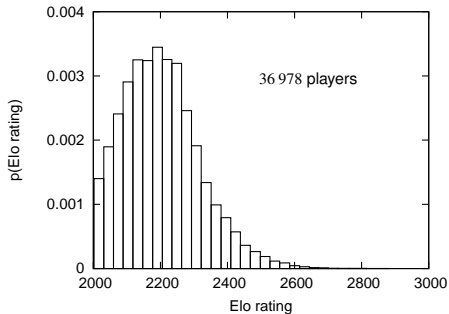
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# Example: chess ratings

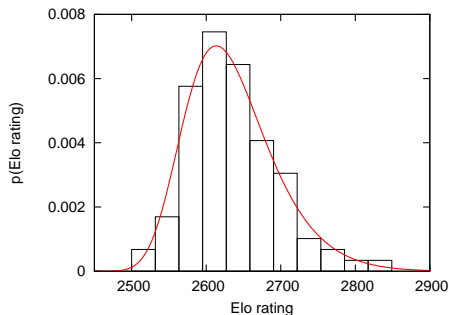
FIDE Elo ratings, January 2001



# Example: chess ratings

Block maxima

$$G_{\xi; \mu, \sigma}(x) = \exp \left[ - \left( 1 + \xi \frac{(x - \mu)}{\sigma} \right)^{-1/\xi} \right].$$

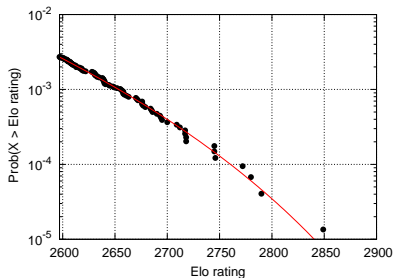
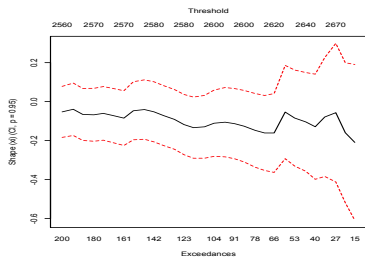


$$\xi = -0.08(6), \mu = 2609(6), \sigma = 53(4)$$

$n = 400$  (block size)

# Example: chess ratings

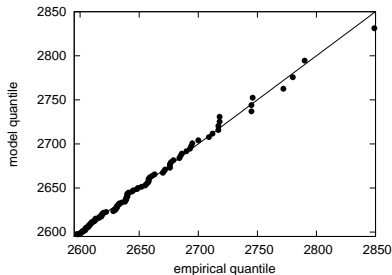
## Exceedances



$$\bar{F}_{\xi; \beta}(y) = \frac{n_u}{n} \left[ 1 + \frac{\xi}{\beta}(y - u) \right]^{-1/\xi}$$

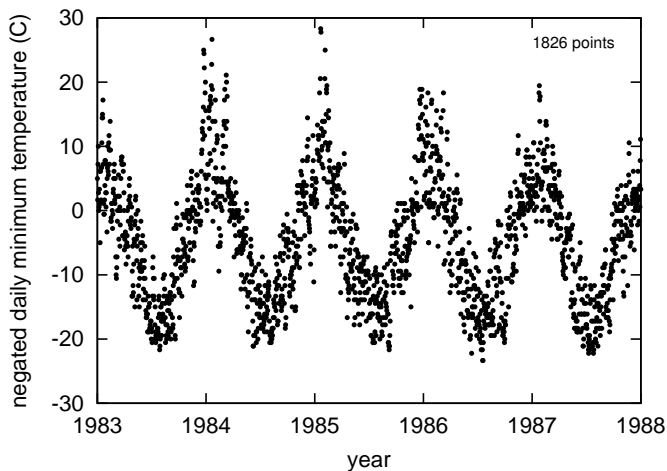
$$\xi = -0.13(9), \beta = 60(8)$$

$$u = 2597 \quad (100 \text{ exceedances})$$

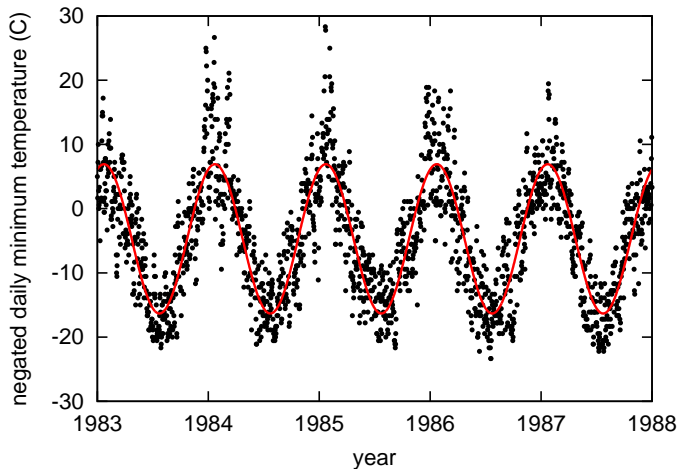


# Example: Wooster daily minimum temperatures

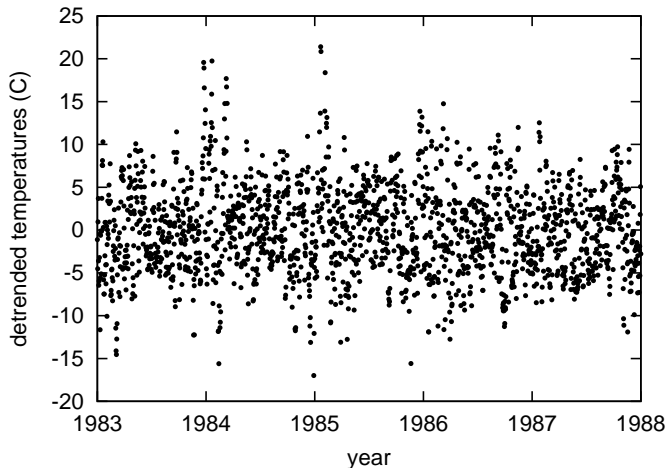
S. G. Coles, J. A. Tawn, R. L. Smith, *Environmetrics* **5**, 221 (1994)



## Example: Wooster daily minimum temperatures

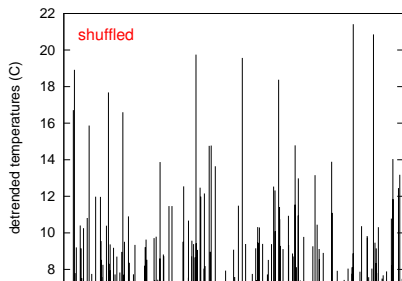
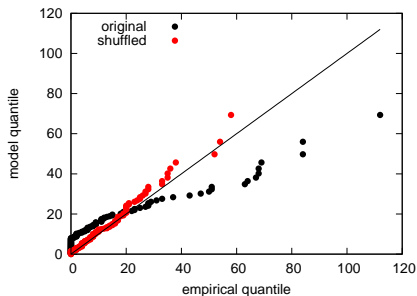
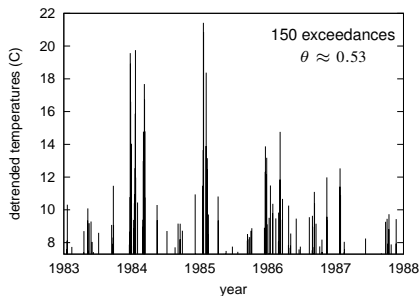


## Example: Wooster daily minimum temperatures



should also detrend the seasonal scale!

# Exceedance arrivals (Poisson for $\theta = 1$ )

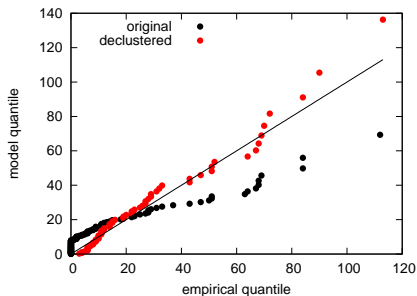
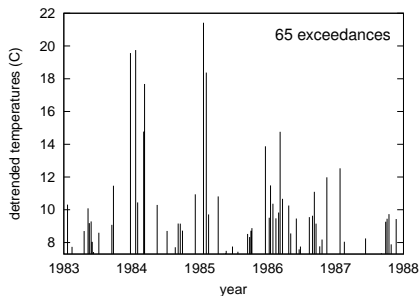


$$\mathbb{P}(\text{interval} > t) = \exp(-\lambda t)$$

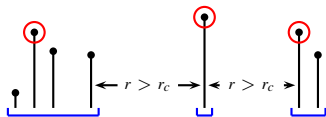


# Declustering exceedances

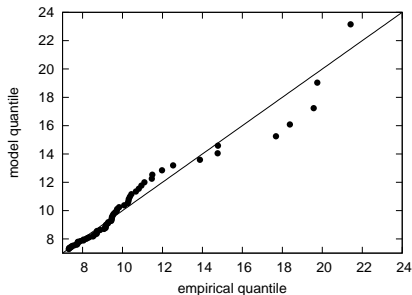
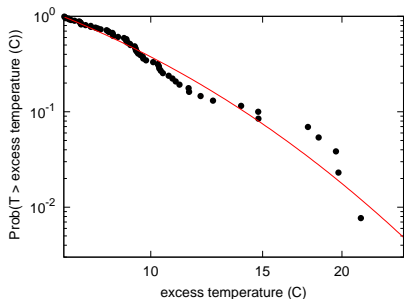
“Runs” method (clusters separated by  $r > r_c$ )



$$\mathbb{P}(\text{interval} > t) = \exp(-\lambda t)$$



# Declustering exceedances

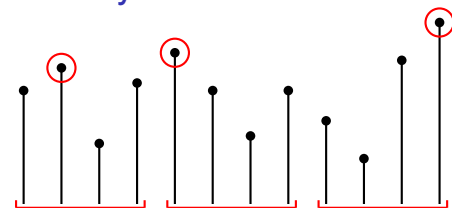


$$\bar{F}_{\xi;\beta}(y) = \frac{n_u}{n} \left[ 1 + \frac{\xi}{\beta}(y - u) \right]^{-1/\xi}$$

$$\xi = 0.08(14), \beta = 2.7(5)$$

(65 declustered exceedances)

# Summary

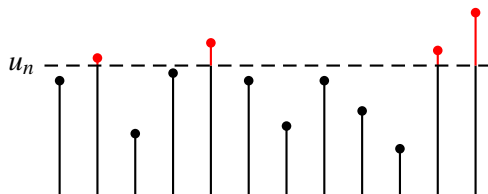


$X_1 X_2 \dots$

$$G_{\xi}(x) = \exp \left\{ - \left[ 1 + \xi \frac{(x-\mu)}{\sigma} \right]^{-1/\xi} \right\}$$

GEV

Also extends to stationary dependent sequences satisfying  $D(u_n)$  conditions!



$X_1 X_2 \dots$

$$H_{\xi}(y) = 1 - \left( 1 + \frac{\xi}{\beta} y \right)^{-1/\xi}$$

GPD